

# How Iris Recognition Works

John Daugman, PhD, OBE

University of Cambridge, The Computer Laboratory, Cambridge CB2 3QG, U.K.  
[www.CL.cam.ac.uk/users/jgd1000/](http://www.CL.cam.ac.uk/users/jgd1000/)

## Abstract

Algorithms developed by the author for recognizing persons by their iris patterns have now been tested in six field and laboratory trials, producing no false matches in several million comparison tests. The recognition principle is the failure of a test of statistical independence on iris phase structure encoded by multi-scale quadrature wavelets. The combinatorial complexity of this phase information across different persons spans about 244 degrees of freedom and generates a discrimination entropy of about 3.2 bits/mm<sup>2</sup> over the iris, enabling real-time decisions about personal identity with extremely high confidence. The high confidence levels are important because they allow very large databases to be searched exhaustively (one-to-many “identification mode”) without making any false matches, despite so many chances. Biometrics lacking this property can only survive one-to-one (“verification”) or few comparisons. This paper explains the algorithms for iris recognition, and presents the results of 2.3 million comparisons among eye images acquired in trials in Britain, the USA, and Japan.

## 1 Introduction

Reliable automatic recognition of persons has long been an attractive goal. As in all pattern recognition problems, the key issue is the relation between inter-class and intra-class variability: objects can be reliably classified only if the variability among different instances of a given class is less than the variability between different classes. For example in face recognition, difficulties arise from the fact that the face is a changeable social organ displaying a variety of expressions, as well as being an active 3D object whose image varies with viewing angle, pose, illumination, accoutrements, and age. It has been shown that for facial images taken at least one year apart, even the best current algorithms have error rates of 43% (Phillips et al. 2000) to 50% (Pentland et al. 2000). Against this intra-class (same face) variability, inter-class variability is limited because different faces possess the same basic set of features, in the same canonical geometry.

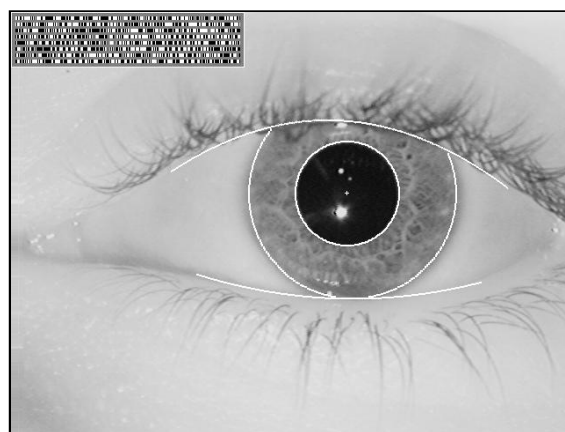


Figure 1: Example of an iris pattern, imaged monochromatically at a distance of about 35 cm. The outline overlay shows results of the iris and pupil localization and eyelid detection steps. The bit stream in the top left is the result of demodulation with complex-valued 2D Gabor wavelets to encode the phase sequence of the iris pattern.

For all of these reasons, iris patterns become interesting as an alternative approach to reliable visual recognition of persons when imaging can be done at distances of less than a meter, and especially when there is a need to search very large databases without incurring any false matches despite a huge number of possibilities. Although small (11 mm) and sometimes problematic to image, the iris has the great mathematical advantage that its pattern variability among different persons is enormous. In addition, as an internal (yet externally visible) organ of the eye, the iris is well protected from the environment, and stable over time. As a planar object its image is relatively insensitive to angle of illumination, and changes in viewing angle cause only affine transformations; even the non-affine pattern distortion caused by pupillary dilation is readily reversible. Finally, the ease of localizing eyes in faces, and the distinctive annular shape of the iris, facilitate reliable and precise isolation of this feature and the creation of a size-invariant representation.

The iris begins to form in the third month of gestation (Kronfeld 1962) and the structures creating its pattern are largely complete by the eighth month, although pigment accretion can continue into the first postnatal years. Its complex pattern can contain many distinctive features such as arching ligaments, furrows, ridges, crypts, rings, corona, freckles, and a zigzag collarette, some of which may be seen in Figure 1. Iris colour is determined mainly by the density of melanin pigment (Chedekel 1994) in its anterior layer and stroma, with blue irises resulting from an absence of pigment: long wavelength light penetrates and is absorbed by the pigment epithelium, while shorter wavelengths are reflected and scattered by the stroma. The striated trabecular meshwork of elastic pectinate ligament creates the predominant texture under visible light, whereas in the near infrared (NIR) wavelengths used for unobtrusive imaging at distances of up to 1 meter, deeper and somewhat more slowly modulated stromal features dominate the iris pattern. In NIR wavelengths, even darkly pigmented irises reveal rich and complex features.

Algorithms described in (Daugman 1993, 1994) for encoding and recognizing iris patterns have been the executable software used in all iris recognition systems so far deployed commercially or in public trials, including those by British Telecom, US Sandia Labs, UK National Physical Laboratory, NCR, Oki, IriScan, Iridian, Sensar, and Sarnoff. All of these organizations have reported a false match rate of 0 in all of their tests, some of which involved millions of different iris pairings. This paper explains how the algorithms work, and presents new data on the statistical properties and singularity of iris patterns based on 2.3 million comparisons.

## 2 Finding an Iris in an Image

To capture the rich details of iris patterns, an imaging system should resolve a minimum of 70 pixels in iris radius. In the field trials to date, a resolved iris radius of 100 to 140 pixels has been more typical. Monochrome CCD cameras (480 x 640) have been used because NIR illumination in the 700nm - 900nm band was required for imaging to be invisible to humans. Some imaging platforms deployed a wide-angle camera for coarse localization of eyes in faces, to steer the optics of a narrow-angle pan/tilt camera that acquired higher resolution images of eyes. There exist many alternative methods for finding and tracking facial features such as the eyes, and this well-researched topic will not be discussed further here. In these trials, most imaging was done without active pan/tilt camera optics, but instead exploited visual feedback via a mirror or video image to enable cooperating Subjects to position their own eyes within the field of view of a single narrow-angle camera.

Focus assessment was performed in real-time (faster than video frame rate) by measuring the total high-frequency power in the 2D Fourier spectrum of each frame, and seeking to maximize this quantity either by moving an active lens or by providing audio feedback to Subjects to adjust their range appropriately. Images passing a minimum focus criterion were then analyzed to find the iris, with precise localization of its boundaries using a coarse-to-fine strategy terminating in single-pixel precision estimates of the center coordinates and radius of both the iris and the pupil. Although the results of the iris search greatly constrain the pupil search, concentricity of these boundaries cannot be assumed. Very often the pupil center is nasal, and inferior, to the iris center. Its radius can range from 0.1 to 0.8 of the iris radius. Thus, all three parameters defining the pupillary circle must be estimated separately from those of the iris. A very effective integrodifferential operator for determining these parameters is:

$$\max_{(r, x_0, y_0)} \left| G_\sigma(r) * \frac{\partial}{\partial r} \oint_{r, x_0, y_0} \frac{I(x, y)}{2\pi r} ds \right| \quad (1)$$

where  $I(x, y)$  is an image such as Fig 1 containing an eye. The operator searches over the image domain  $(x, y)$  for the maximum in the blurred partial derivative with respect to increasing radius  $r$ , of the normalized contour integral of  $I(x, y)$  along a circular arc  $ds$  of radius  $r$  and center coordinates  $(x_0, y_0)$ . The symbol  $*$  denotes convolution and  $G_\sigma(r)$  is a smoothing function such as a Gaussian of scale  $\sigma$ . The complete operator behaves in effect as a circular edge detector, blurred at a scale set by  $\sigma$ , which searches iteratively for a maximum contour integral derivative with increasing radius at successively finer scales of analysis through the three parameter space of center coordinates and radius  $(x_0, y_0, r)$  defining a path of contour integration.

The operator in (1) serves to find both the pupillary boundary and the outer (limbus) boundary of the iris, although the initial search for the limbus also incorporates evidence of an interior pupil to improve its robustness since the limbic boundary itself usually has extremely soft contrast when long wavelength NIR illumination is used. Once the coarse-to-fine iterative searches for both these boundaries have reached single pixel precision, then a similar approach to detecting curvilinear edges is used to localize both the upper and lower eyelid boundaries. The path of contour integration in (1) is changed from circular to arcuate, with spline parameters fitted by standard statistical estimation methods to describe optimally the available evidence for each eyelid boundary. The result of all these localization operations is the isolation of iris tissue from other image regions, as illustrated in Fig 1 by the graphical overlay on the eye.

### Phase-Quadrant Demodulation Code

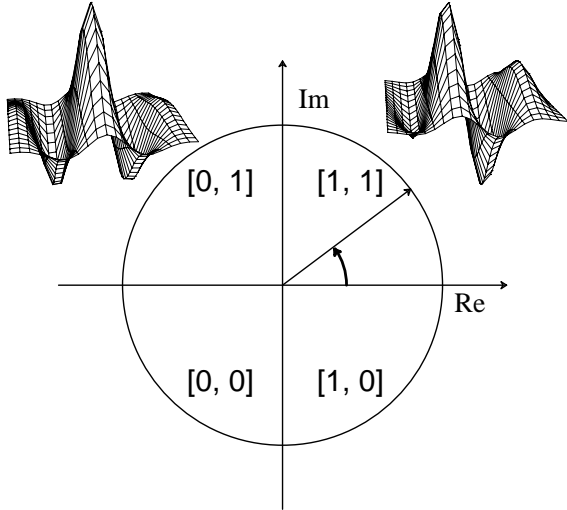


Figure 2: The phase demodulation process used to encode iris patterns. Local regions of an iris are projected (Eq 2) onto quadrature 2D Gabor wavelets, generating complex-valued coefficients whose real and imaginary parts specify the coordinates of a phasor in the complex plane. The angle of each phasor is quantized to one of the four quadrants, setting two bits of phase information. This process is repeated all across the iris with many wavelet sizes, frequencies, and orientations, to extract 2,048 bits.

### 3 Iris Feature Encoding by 2D Wavelet Demodulation

Each isolated iris pattern is then demodulated to extract its phase information using quadrature 2D Gabor wavelets (Daugman 1985, 1988, 1994). This encoding process is illustrated in Fig 2. It amounts to a patch-wise phase quantization of the iris pattern, by identifying in which quadrant of the complex plane each resultant phasor lies when a given area of the iris is projected onto complex-valued 2D Gabor wavelets:

$$h_{\{Re, Im\}} = \text{sgn}_{\{Re, Im\}} \int_{\rho} \int_{\phi} I(\rho, \phi) e^{-i\omega(\theta_0 - \phi)} \cdot e^{-(r_0 - \rho)^2 / \alpha^2} e^{-(\theta_0 - \phi)^2 / \beta^2} \rho d\rho d\phi \quad (2)$$

where  $h_{\{Re, Im\}}$  can be regarded as a complex-valued bit whose real and imaginary parts are either 1 or 0 (sgn) depending on the sign of the 2D integral;  $I(\rho, \phi)$  is the raw iris image in a dimensionless polar coordinate system that is size- and translation-invariant, and which also corrects for pupil dilation as explained in a later section;  $\alpha$  and  $\beta$  are the multi-scale 2D wavelet size parameters, spanning an 8-fold range from 0.15mm to 1.2mm on the iris;  $\omega$  is wavelet frequency, spanning 3 octaves in inverse proportion to  $\beta$ ; and  $(r_0, \theta_0)$  represent the polar coordinates of each region of iris for which the phasor coordinates  $h_{\{Re, Im\}}$

are computed. Such a phase quadrant coding sequence is illustrated for one iris by the bit stream shown graphically in Fig 1. A desirable feature of the phase code portrayed in Fig 2 is that it is a cyclic, or grey code: in rotating between any adjacent phase quadrants, only a single bit changes, unlike a binary code in which two bits may change, making some errors arbitrarily more costly than others. Altogether 2,048 such phase bits (256 bytes) are computed for each iris, but in a major improvement over the earlier (Daugman 1993) algorithms, now an equal number of masking bits are also computed to signify whether any iris region is obscured by eyelids, contains any eyelash occlusions, specular reflections, boundary artifacts of hard contact lenses, or poor signal-to-noise ratio and thus should be ignored in the demodulation code as artifact.

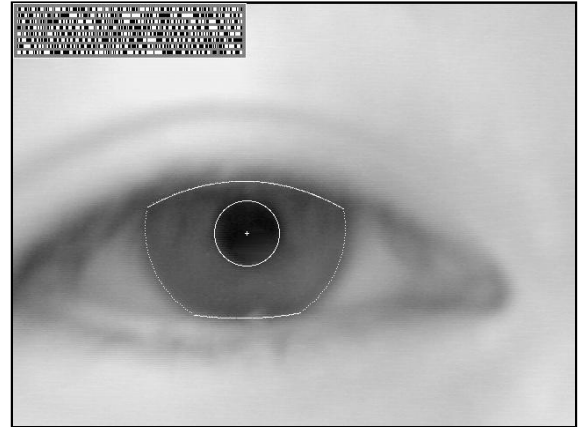


Figure 3: Illustration that even for poorly focused eye images, the bits of a demodulation phase sequence are still set, primarily by random CCD noise. This prevents poorly focused eye images from resembling each other in the pattern matching stage, in the way that (e.g.) poorly resolved face images look alike and can be confused with each other.

Only phase information is used for recognizing irises because amplitude information is not very discriminating, and it depends upon extraneous factors such as imaging contrast, illumination, and camera gain. The phase bit settings which code the sequence of projection quadrants as shown in Fig 2 capture the information of wavelet zero-crossings, as is clear from the sign operator in (2). The extraction of phase has the further advantage that phase angles are assigned regardless of how low the image contrast may be, as illustrated by the extremely out-of-focus image in Fig 3. Its phase bit stream has statistical properties such as run lengths similar to those of the code for the properly focused eye image in Fig 1. (Fig 3 also illustrates the robustness of the iris- and pupil-finding operators, and the eyelid detection operators, despite poor focus.) The benefit which arises from

the fact that phase bits are set also for a poorly focused image as shown here, even if based only on random CCD noise, is that different poorly focused irises never become confused with each other when their phase codes are compared. By contrast, images of different faces look increasingly alike when poorly resolved, and may be confused with each other by appearance-based face recognition algorithms.

#### 4 The Test of Statistical Independence: Combinatorics of Phase Sequences

The key to iris recognition is the failure of a test of statistical independence, which involves so many degrees-of-freedom that this test is virtually guaranteed to be passed whenever the phase codes for two different eyes are compared, but to be uniquely failed when any eye’s phase code is compared with another version of itself.

The test of statistical independence is implemented by the simple Boolean Exclusive-OR operator (XOR) applied to the 2,048 bit phase vectors that encode any two iris patterns, masked (AND’ed) by both of their corresponding mask bit vectors to prevent non-iris artifacts from influencing iris comparisons. The XOR operator  $\otimes$  detects disagreement between any corresponding pair of bits, while the AND operator  $\cap$  ensures that the compared bits are both deemed to have been uncorrupted by eyelashes, eyelids, specular reflections, or other noise. The norms ( $\| \cdot \|$ ) of the resultant bit vector and of the AND’ed mask vectors are then measured in order to compute a fractional Hamming Distance (HD) as the measure of the dissimilarity between any two irises, whose two phase code bit vectors are denoted  $\{codeA, codeB\}$  and whose mask bit vectors are denoted  $\{maskA, maskB\}$ :

$$HD = \frac{\|(codeA \otimes codeB) \cap maskA \cap maskB\|}{\|maskA \cap maskB\|} \quad (3)$$

The denominator tallies the total number of phase bits that mattered in iris comparisons after artifacts such as eyelashes and specular reflections were discounted, so the resulting HD is a fractional measure of dissimilarity; 0 would represent a perfect match. The Boolean operators  $\otimes$  and  $\cap$  are applied in vector form to binary strings of length up to the word length of the CPU, as a single machine instruction. Thus for example on an ordinary 32-bit machine, any two integers between 0 and 4 billion can be XOR’ed in a single machine instruction to generate a third such integer, each of whose bits in a binary expansion is the XOR of the corresponding pair of bits of the original two integers. This implementation of (3) in parallel 32-bit chunks enables extremely rapid comparisons of iris codes when searching through a large database to find a match. On a 300 MHz CPU, such exhaus-

Binomial Distribution of IrisCode Hamming Distances

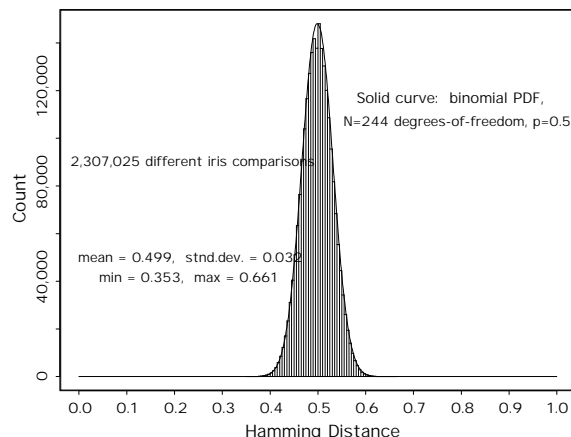


Figure 4: Distribution of Hamming Distances from all 2.3 million possible comparisons between different pairs of irises in the database. The histogram forms a perfect binomial distribution with  $p = 0.5$  and  $N = 244$  degrees-of-freedom, as shown by the solid curve (Eq 4). The data implies that it is extremely improbable for two different irises to disagree in less than about a third of their phase information.

tive searches are performed at a rate of about 100,000 irises per second.

Because any given bit in the phase code for an iris is equally likely to be 1 or 0, and because different irises are uncorrelated, the expected proportion of agreeing bits between the codes for two different irises is  $HD = 0.500$ . The histogram in Fig 4 shows the distribution of HDs obtained from 2.3 million comparisons between different pairings of iris images acquired by licensees of these algorithms in the UK, the USA, and Japan. There were 2,150 different iris images, including 10 each of one subset of 70 eyes. Excluding those duplicates of  $(700 \times 9)$  same-eye comparisons, and not double-counting pairs, and not comparing any image with itself, the total number of unique pairings between different eye images whose HDs could be computed was  $((2,150 \times 2,149 - 700 \times 9) / 2) = 2,307,025$ . Their observed mean HD was  $p = 0.499$  with standard deviation  $\sigma = 0.032$ ; their full distribution in Fig 4 corresponds to a binomial having  $N = p(1 - p) / \sigma^2 = 244$  degrees-of-freedom, as shown by the solid curve. The extremely close fit of the theoretical binomial to the observed distribution is a consequence of the fact that each comparison between two phase code bits from two different irises is essentially a Bernoulli trial, albeit with correlations between successive “coin tosses.”

In the phase code for any given iris, only small subsets of bits are mutually independent due to the internal correlations, especially radial, within an iris. (If all  $N = 2,048$  phase bits were independent, then the distribution in Fig 4 would be very much sharper, with an expected standard deviation of only

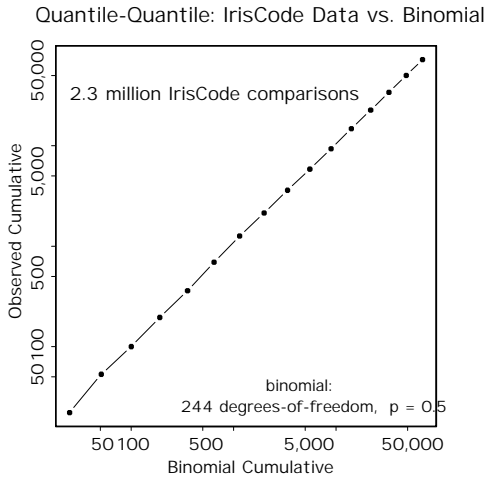


Figure 5: Quantile-quantile plot of the observed cumulatives under the left tail of the histogram in Fig 4, versus the predicted binomial cumulatives. Their close agreement over several orders of magnitude strongly confirms the binomial model for phase bit comparisons between different irises.

$\sqrt{p(1-p)/N} = 0.011$  and so the HD interval between 0.49 and 0.51 would contain most of the distribution.) Bernoulli trials that are correlated (Viveros et al. 1984) remain binomially distributed but with a reduction in  $N$ , the effective number of tosses, and hence an increase in the  $\sigma$  of the normalized HD distribution. The form and width of the HD distribution in Fig 4 tell us that the amount of difference between the phase codes for different irises is distributed equivalently to runs of 244 tosses of a fair coin (Bernoulli trials with  $p = 0.5$ ,  $N = 244$ ). Expressing this variation as a discrimination entropy (Cover and Thomas 1991) and using typical iris and pupil diameters of 11mm and 5mm respectively, the observed amount of statistical variability among different iris patterns corresponds to an information density of about 3.2 bits/mm<sup>2</sup> on the iris.

The theoretical binomial distribution plotted as the solid curve in Fig 4 has the fractional functional form

$$f(x) = \frac{N!}{m!(N-m)!} p^m (1-p)^{(N-m)} \quad (4)$$

where  $N = 244$ ,  $p = 0.5$ , and  $x = m/N$  is the outcome fraction of  $N$  Bernoulli trials (e.g. coin tosses that are “heads” in each run). In our case,  $x$  is the HD, the fraction of phase bits that happen to agree when two different irises are compared. To validate such a statistical model we must also study the behaviour of the tails, by examining quantile-quantile plots of the observed cumulatives versus the theoretically predicted cumulatives from 0 up to sequential points in the tail. Such a “Q-Q” plot is given in Fig 5. The straight line relationship reveals very precise agreement between model and data, over a range of more than three orders of magnitude. It is clear from both Figures 4 and 5 that it is extremely improbable that

Genetically Identical Eyes Have Uncorrelated IrisCodes

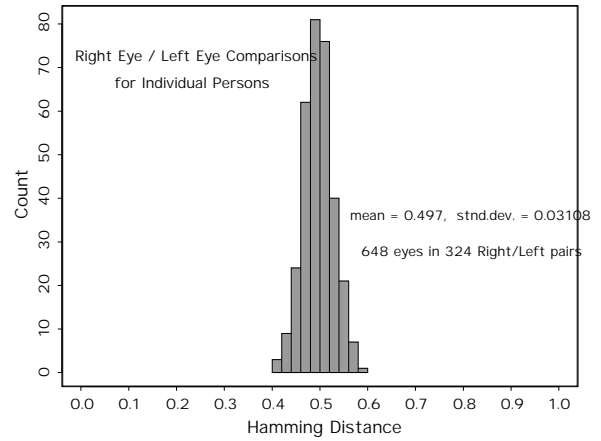


Figure 6: Distribution of Hamming Distances between genetically identical irises, in 648 paired eyes from 324 persons. The data are statistically indistinguishable from that shown in Fig 4 comparing unrelated irises. Unlike eye colour, the phase structure of iris patterns therefore appears to be epigenetic, arising from random events and circumstances in the morphogenesis of this tissue.

two different irises might disagree by chance in fewer than at least a third of their bits. (Of the 2.3 million iris comparisons plotted in the histogram of Figure 4, the smallest Hamming Distance observed was 0.353.) Computing the cumulative of  $f(x)$  from 0 to 0.333 indicates that the probability of such an event is about 1 in 10 million. The cumulative from 0 to just 0.300 is 1 in 7 billion. Thus, even the observation of a relatively poor degree of match between the phase codes for two different iris images (say, 70% agreement or HD = 0.300) would still provide extraordinarily compelling evidence of identity, because the test of statistical independence is still failed so convincingly.

I also compared genetically identical eyes in the same manner, in order to discover the degree to which their textural patterns were correlated and hence genetically determined. A convenient source of genetically identical irises are the right and left pair from any given person; such pairs have the same genetic relationship as the four irises of monozygotic twins, or indeed the prospective  $2N$  irises of  $N$  clones. Although eye colour is of course strongly determined genetically, as is overall iris appearance, the detailed patterns of genetically identical irises appear to be as uncorrelated as they are among unrelated eyes. Using the same methods as described above, 648 right/left iris pairs from 324 persons were compared pairwise. Their mean HD was 0.497 with standard deviation 0.031, and their distribution (Fig 6) was statistically indistinguishable from the distribution for unrelated eyes (Fig 4). A set of 6 pairwise comparisons among the eyes of actual monozygotic twins also yielded a result (mean HD = 0.507) expected for unrelated eyes.

It appears that the phenotypic random patterns visible in the human iris are almost entirely epigenetic.

## 5 Recognizing Irises Regardless of Size, Position, and Orientation

Robust representations for pattern recognition must be invariant to changes in the size, position, and orientation of the patterns. In the case of iris recognition, this means we must create a representation that is invariant to the optical size of the iris in the image (which depends upon the distance to the eye, and the camera optical magnification factor); the size of the pupil within the iris (which introduces a non-affine pattern deformation); the location of the iris within the image; and the iris orientation, which depends upon head tilt, torsional eye rotation within its socket (cyclovergence), and camera angles, compounded with imaging through pan/tilt eye-finding mirrors that introduce additional image rotation factors as a function of eye position, camera position, and mirror angles. Fortunately, invariance to all of these factors can readily be achieved.

For on-axis but possibly rotated iris images, it is natural to use a projected pseudo polar coordinate system. The polar coordinate grid is not necessarily concentric, since in most eyes the pupil is not central in the iris; it is not unusual for its nasal displacement to be as much as 15%. This coordinate system can be described as doubly-dimensionless: the polar variable, angle, is inherently dimensionless, but in this case the radial variable is also dimensionless, because it ranges from the pupillary boundary to the limbus always as a unit interval  $[0, 1]$ . The dilation and constriction of the elastic meshwork of the iris when the pupil changes size is intrinsically modelled by this coordinate system as the stretching of a homogeneous rubber sheet, having the topology of an annulus anchored along its outer perimeter, with tension controlled by an (off-centered) interior ring of variable radius.

The homogeneous rubber sheet model assigns to each point on the iris, regardless of its size and pupillary dilation, a pair of real coordinates  $(r, \theta)$  where  $r$  is on the unit interval  $[0, 1]$  and  $\theta$  is angle  $[0, 2\pi]$ . The remapping of the iris image  $I(x, y)$  from raw cartesian coordinates  $(x, y)$  to the dimensionless non-concentric polar coordinate system  $(r, \theta)$  can be represented as

$$I(x(r, \theta), y(r, \theta)) \rightarrow I(r, \theta) \quad (5)$$

where  $x(r, \theta)$  and  $y(r, \theta)$  are defined as linear combinations of both the set of pupillary boundary points  $(x_p(\theta), y_p(\theta))$  and the set of limbus boundary points along the outer perimeter of the iris  $(x_s(\theta), y_s(\theta))$  bordering the sclera, both of which are detected by

finding the maximum of the operator (1).

$$x(r, \theta) = (1 - r)x_p(\theta) + rx_s(\theta) \quad (6)$$

$$y(r, \theta) = (1 - r)y_p(\theta) + ry_s(\theta) \quad (7)$$

Since the radial coordinate ranges from the iris inner boundary to its outer boundary as a unit interval, it inherently corrects for the elastic pattern deformation in the iris when the pupil changes in size.

The localization of the iris and the coordinate system described above achieve invariance to the 2D position and size of the iris, and to the dilation of the pupil within the iris. However, it would not be invariant to the orientation of the iris within the image plane. The most efficient way to achieve iris recognition with orientation invariance is not to rotate the image itself using the Euler matrix, but rather to compute the iris phase code in a single canonical orientation and then to compare this very compact representation at many discrete orientations by cyclic scrolling of its angular variable. The statistical consequences of seeking the best match after numerous relative rotations of two iris codes are straightforward. Let  $f_0(x)$  be the raw density distribution obtained for the HDs between different irises after comparing them only in a single relative orientation; for example,  $f_0(x)$  might be the binomial defined in (4). Then  $F_0(x)$ , the cumulative of  $f_0(x)$  from 0 to  $x$ , becomes the probability of getting a false match in such a test when using HD acceptance criterion  $x$ :

$$F_0(x) = \int_0^x f_0(x) dx \quad (8)$$

or, equivalently,

$$f_0(x) = \frac{d}{dx} F_0(x) \quad (9)$$

Clearly, then, the probability of not making a false match when using criterion  $x$  is  $1 - F_0(x)$  after a single test, and it is  $[1 - F_0(x)]^n$  after carrying out  $n$  such tests independently at  $n$  different relative orientations. It follows that the probability of a false match after a “best of  $n$ ” test of agreement, when using HD criterion  $x$ , regardless of the actual form of the raw unrotated distribution  $f_0(x)$ , is:

$$F_n(x) = 1 - [1 - F_0(x)]^n \quad (10)$$

and the expected density  $f_n(x)$  associated with this cumulative is

$$\begin{aligned} f_n(x) &= \frac{d}{dx} F_n(x) \\ &= n f_0(x) [1 - F_0(x)]^{n-1} \end{aligned} \quad (11)$$

Each of the 2.3 million pairings of different iris images whose HD distribution was shown in Fig 4,

IrisCode Comparisons after Rotations: Best Matches

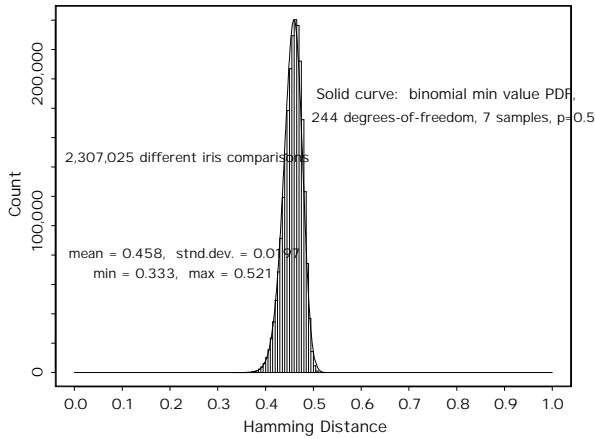


Figure 7: Distribution of Hamming Distances from the same set of 2.3 million comparisons as seen in Fig 4, but allowing for 7 relative rotations and preserving only the best match found for each pair. This “best of  $n$ ” test skews the distribution to the left and reduces its mean from about 0.5 to 0.458. The solid curve is the theoretical prediction for such “extreme-value” sampling, as described by Eqts (4) and (8) - (11).

was submitted to further comparisons in each of 7 relative orientations. This generated 16.1 million HD outcomes, but in each group of 7 associated with any one pair of irises, only the best match (smallest HD) was retained. The histogram of these new 2.3 million best HDs is shown in Fig 7. Since only the smallest value in each group of 7 samples was retained, the new distribution is skewed and biased to a lower mean value (HD = 0.458), as expected from the theory of extreme value sampling. The solid curve in Fig 7 is a plot of (11), incorporating (4) and (8) as its terms, and it shows an excellent fit between theory (binomial extreme value sampling) and data. The fact that the minimum HD observed in all of these rotated comparisons was never smaller than 0.333 illustrates the extreme improbability that the phase sequences for two different irises might disagree in fewer than a third of their bits. This suggests that in order to identify people by their iris patterns with high confidence, we need to demand only a very forgiving degree of match (say, HD < 0.330).

## 6 Uniqueness of Failing the Test of Statistical Independence

The statistical data and theory presented above show that we can perform iris recognition successfully just by a test of statistical independence. Any two different irises are statistically “guaranteed” to pass this test of independence, and any two images that fail this test (i.e. produce HD < 0.330) must be images of the same iris. Thus, it is the unique failure of the test of independence, that is the basis for iris recognition.

False Match Probabilities: Cumulatives under PDF

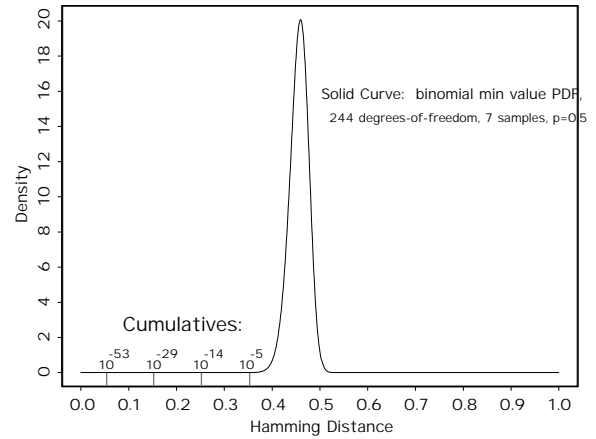


Figure 8: Calculated cumulatives under the left tail of the distribution seen in Fig 7, up to sequential points, using the functional analysis described by Eqts (4) and (8) - (11). The extremely rapid attenuation of these cumulatives reflects the binomial combinatorics that dominate Eq (4). This accounts for the astronomical confidence levels against a false match, when executing this test of statistical independence.

It is informative to calculate the significance of any observed HD matching score, in terms of the likelihood that it could have arisen by chance from two different irises. These probabilities give a confidence level associated with any recognition decision. Fig 8 shows the false match probabilities marked off in cumulatives along the tail of the distribution presented in Fig 7 (same theoretical curve (11) as plotted in Fig 7 and with the justification presented in Fig 4 and Fig 5.) Table 1 enumerates the cumulatives of (11) (false match probabilities) as a more fine-grained function of HD decision criterion in the range between 0.26 and 0.35. Calculation of the large factorial terms in (4) was done with Stirling’s approximation which errs by less than 1% for  $n \geq 9$ :

$$n! \approx \exp(n \ln(n) - n + \frac{1}{2} \ln(2\pi n)) \quad (12)$$

HD Criterion	Odds of False Match
0.26	1 in $10^{13}$
0.27	1 in $10^{12}$
0.28	1 in 84 billion
0.29	1 in 8.6 billion
0.30	1 in 1 billion
0.31	1 in 127 million
0.32	1 in 18 million
0.33	1 in 2.9 million
0.34	1 in 527,000
0.35	1 in 105,000

Table 1: Cumulatives under (11) giving single false match probabilities for various HD criteria.

The practical importance of the astronomical odds against a false match when the match quality is better than about  $HD \leq 0.33$ , as shown in Fig 8 and in Table 1, is that such high confidence levels allow very large databases to be searched exhaustively without succumbing to any of the many opportunities for suffering a false match. The requirements of operating in one-to-many “identification” mode are vastly more demanding than operating merely in one-to-one “verification” mode (in which an identity must first be explicitly asserted, which is then verified in a yes/no decision by comparison against just the single nominated template).

If  $P_1$  is the false match probability for single one-to-one verification trials, then clearly  $P_N$ , the probability of making at least one false match when searching a database of  $N$  unrelated patterns, is:

$$P_N = 1 - (1 - P_1)^N \quad (13)$$

because  $(1 - P_1)$  is the probability of not making a false match in single comparisons; this must happen  $N$  independent times; and so  $(1 - P_1)^N$  is the probability that such a false match never occurs.

It is interesting to consider how a seemingly impressive biometric one-to-one “verifier” would perform in exhaustive search mode once databases become larger than about 100, in view of (13). For example, a face recognition algorithm that truly achieved 99.9% correct rejection when tested on non-identical faces, hence making only 0.1% false matches, would seem to be performing at a very impressive level because it must confuse no more than 10% of all identical twin pairs (since about 1% of all persons in the general population have an identical twin). But even with its  $P_1 = 0.001$ , how good would it be for searching large databases?

Using (13) we see that when the search database size has reached merely  $N = 200$  unrelated faces, the probability of at least one false match among them is already 18%. When the search database is just  $N = 2000$  unrelated faces, the probability of at least one false match has reached 86%. Clearly, identification is vastly more demanding than one-to-one verification, and even for moderate database sizes, merely “good” verifiers are of no use as identifiers. Observing the approximation that  $P_N \approx NP_1$  for small  $P_1 \ll \frac{1}{N} \ll 1$ , when searching a database of size  $N$  an identifier needs to be roughly  $N$  times better than a verifier to achieve comparable odds against making false matches.

The algorithms for iris recognition exploit the extremely rapid attenuation of the HD distribution tail created by binomial combinatorics, to accommodate very large database searches without suffering false

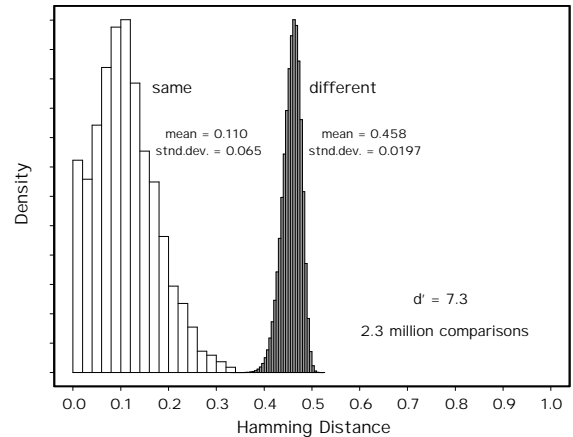


Figure 9: The Decision Environment for iris recognition under relatively unfavourable conditions, using images acquired at different distances, and by different optical platforms.

matches. The HD threshold is adaptive, to maintain  $P_N < 10^{-6}$  regardless of how large the search database size  $N$  is. As Table 1 illustrates, this means that if the search database contains 1 million different iris patterns, it is only necessary for the HD match criterion to adjust downwards from 0.33 to 0.27 in order to maintain still a net false match probability of  $10^{-6}$  for the entire database.

## 7 Decision Environment for Iris Recognition

The overall “decidability” of the task of recognizing persons by their iris patterns is revealed by comparing the Hamming Distance distributions for same versus for different irises. The left distribution in Fig 9 shows the HDs computed between 7,070 different pairs of same-eye images at different times, under different conditions, and usually with different cameras; and the right distribution gives the same 2.3 million comparisons among different eyes shown earlier. To the degree that one can confidently decide whether an observed sample belongs to the left or the right distribution in Fig 9, iris recognition can be successfully performed. Such a dual distribution representation of the decision problem may be called the “decision environment,” because it reveals the extent to which the two cases (same versus different) are separable and thus how reliably decisions can be made, since the overlap between the two distributions determines the error rates.

Whereas Fig 9 shows the decision environment under less favourable conditions (images acquired by different camera platforms), Fig 10 shows the decision environment under ideal (almost artificial) conditions. Subjects’ eyes were imaged in a laboratory setting using always the same camera with fixed zoom factor and at fixed distance, and with fixed illumina-



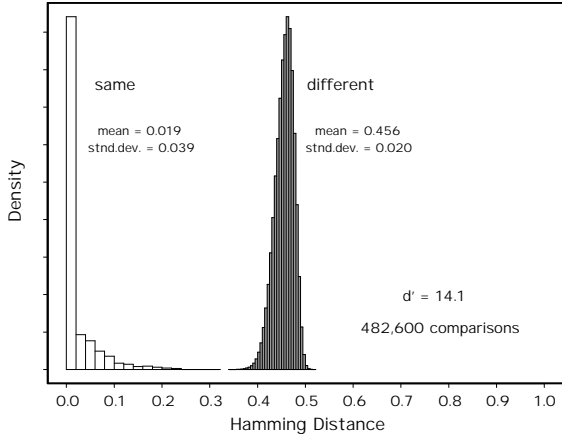


Figure 10: The Decision Environment for iris recognition under very favourable conditions, using always the same camera, distance, and lighting.

tion. Not surprisingly, more than half of such image comparisons achieved an HD of 0.00, and the average HD was a mere 0.019. It is clear from comparing Fig 9 and Fig 10 that the “authentics” distribution for iris recognition (the similarity between different images of the same eye, as shown in the left-side distributions), depends very strongly upon the image acquisition conditions. However, the measured similarity for “imposters” (the right-side distribution) is apparently almost completely independent of imaging factors. Instead, it mainly reflects just the combinatorics of Bernoulli trials, as bits from independent binary sources (the phase codes for different irises) are compared.

For two-choice decision tasks (e.g. same versus different), such as biometric decision making, the “decidability” index  $d'$  measures how well separated the two distributions are, since recognition errors would be caused by their overlap. If their two means are  $\mu_1$  and  $\mu_2$ , and their two standard deviations are  $\sigma_1$  and  $\sigma_2$ , then  $d'$  is defined as

$$d' = \frac{|\mu_1 - \mu_2|}{\sqrt{(\sigma_1^2 + \sigma_2^2)}/2} \quad (14)$$

This measure of decidability is independent of how liberal or conservative is the acceptance threshold used. Rather, by measuring separation, it reflects the degree to which any improvement in (say) the false match error rate must be paid for by a worsening of the failure-to-match error rate. The performance of every biometric technology can be calibrated by its  $d'$  score. The measured decidability for iris recognition is  $d' = 7.3$  for the non-ideal (crossed platform) conditions presented in Fig 9, and it is  $d' = 14.1$  for the ideal imaging conditions presented in Fig 10.

Based on the left-side distributions in Figs 9 and 10, one could calculate a table of probabilities of failure to match, as a function of HD match criterion, just

as we did earlier in Table 1 for false match probabilities based on the right-side distribution. However, such estimates may not be stable because the “authentics” distributions depend strongly on the quality of imaging (e.g. motion blur, focus, noise, etc.) and would be different for different optical platforms. As illustrated earlier by the badly defocused image of Fig 3, phase bits are still set randomly with binomial statistics in poor imaging, and so the right distribution is the stable asymptotic form both in the case of well imaged irises (Fig 10) and poorly imaged irises (Fig 9). Imaging quality determines how much the same-iris distribution evolves and migrates leftward, away from the asymptotic different-iris distribution on the right. In any case, we note that for the 7,070 same-iris comparisons shown in Fig 9, their highest HD was 0.327 which is below the smallest HD of 0.333 for 2.3 million comparisons between different irises. Thus a decision criterion around 0.330 for the empirical data sets shown produces a correct match rate of 100% and no false matches. At this criterion, using the cumulatives of (11) as tabulated in Table 1, the theoretical false match probability is 1 in 2.9 million.

Notwithstanding this diversity among iris patterns and their apparent singularity because of so many dimensions of random variation, their utility as a basis for automatic personal identification would depend upon their relative stability over time. There is a popular belief that the iris changes systematically with one’s health or personality, and even that its detailed features reveal the states of individual organs (“iridology”); but such claims have been discredited (e.g. Berggren 1985; Simon et al. 1979) as medical fraud. In any case, the recognition principle described here is intrinsically tolerant of a large proportion of the iris information being corrupted, say up to about a third, without significantly impairing the inference of personal identity by the simple test of statistical independence.

## 8 Speed Performance Summary

On a 300 MHz Sun workstation, the execution times for the critical steps in iris recognition are as follows, using optimized integer code:

<i>Operation</i>	<i>Time</i>
Assess image focus	15 msec
Scrub specular reflections	56 msec
Localize eye and iris	90 msec
Fit pupillary boundary	12 msec
Detect and fit both eyelids	93 msec
Remove lashes and contact lens edges	78 msec
Demodulation and IrisCode creation	102 msec
XOR comparison of two IrisCodes	10 $\mu$ s

Table 2: Speeds of various stages in the iris recognition process.

The search engine can perform about 100,000 full comparisons between different irises per second, because of the efficient implementation of the matching process in terms of elementary Boolean operators  $\otimes$  and  $\cap$  acting in parallel on the computed phase bit sequences. If database size was measured in millions of enrolled persons, then the inherent parallelism of the search process should be exploited for the sake of speed by dividing up the entire search database into units of about 100,000 persons each. The confidence levels shown in Table 1 indicate how the decision threshold should be adapted for each of these parallel search engines, in order to ensure that no false matches were made despite several large-scale searches being conducted independently. The mathematics of the iris recognition algorithms make it clear that databases the size of entire nations could be searched in parallel to make a confident identification decision, in about 1 second using parallel banks of inexpensive CPUs, if such large national iris databases ever came to exist.

## Bibliography

- Adini, Y., Moses, Y., and Ullman, S. 1997. Face recognition: the problem of compensating for changes in illumination direction. *Trans. Pat. Anal. Mach. Intell.* 19(7): 721-732.
- Belhumeur, P.N., Hespanha, J.P., and Kriegman, D.J. 1997. Eigenfaces vs. Fisherfaces: Recognition using class-specific linear projection. *Trans. Pat. Anal. Mach. Intell.* 19(7): 711-720.
- Berggren, L. 1985. Iridology: A critical review. *Acta Ophthalmologica* 63(1): 1-8.
- Chedekel, M.R. 1995. Photophysics and photochemistry of melanin. In: *Melanin: Its Role in Human Photoprotection*. Valdenmar: Overland Park, 11-23.
- Cover, T. and Thomas, J. 1991. *Elements of Information Theory*. Wiley: New York.
- Daugman, J. 1985. Uncertainty relation for resolution in space, spatial frequency, and orientation optimized by two-dimensional visual cortical filters. *Journal of the Optical Society of America A* 2(7): 1160-1169.
- Daugman, J. 1988. Complete discrete 2D Gabor transforms by neural networks for image analysis and compression. *Trans. Acous. Sp. Sig. Proc.* 36(7): 1169-1179.
- Daugman, J. 1993. High confidence visual recognition of persons by a test of statistical independence. *Trans. Pattern Analysis and Machine Intelligence* 15(11): 1148-1161.
- Daugman, J. 1994. U.S. Patent No. 5,291,560: *Biometric Personal Identification System Based on Iris Analysis*. Issue Date: 1 March 1994.
- Daugman, J., and Downing, C. 1995. Demodulation, predictive coding, and spatial vision. *Journal of the Optical Society of America A* 12(4): 641-660.
- Kronfeld, P. 1962. Gross anatomy and embryology of the eye. In: *The Eye* (H. Davson, Ed.) Academic Press: London.
- Pentland, A., and Choudhury, T. 2000. Face recognition for smart environments. *Computer* 33(2): 50-55.
- Phillips, P.J., Martin, A., Wilson, C.L., and Przybocki, M. 2000. An introduction to evaluating biometric systems. *Computer* 33(2): 56-63.
- Phillips, P.J., Moon, H., Rizvi, S.A., and Rauss, P.J. 2000. The FERET evaluation methodology for face-recognition algorithms. *Trans. Pat. Anal. Mach. Intell.* 22(10): 1090-1104.
- Simon, A., Worthen, D.M., and Mitas, J.A. 1979. An evaluation of iridology. *Journal of the American Medical Association* 242: 1385-1387.
- Viveros, R., Balasubramanian, K., and Balakrishnan, N. 1984. Binomial and negative binomial analogues under correlated Bernoulli trials. *The American Statistician* 48(3): 243-247.